

Chapter 6: Linear Model Selection & Regularization

In the regression setting, the standard linear model is commonly used to describe the relationship between a response Y and a set of variables X_1, \dots, X_p .

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \varepsilon$$

typically fit model using least squares.

↓
we will talk about
other ways we could approach this
fitting problem.

(later will go
non-linear).

The linear model has distinct advantages in terms of inference and is often surprisingly competitive for prediction. How can it be improved?

replace least squares with alternative fitting procedures.

We can yield both better prediction accuracy and model interpretability:

- prediction accuracy: If true relationship is \approx linear, least squares will have low bias.
 - If $n \gg p \Rightarrow$ also low variance \Rightarrow perform well on test data!
 - But n not much larger than $p \Rightarrow$ high variability \Rightarrow poor performance on test data.
 - If $n < p \Rightarrow$ no longer have unique solution \Rightarrow variance $= \infty \Rightarrow$ cannot use this at all!

goal: reduce variance without adding too much bias.
- model interpretability: often many variables used in regression are not in fact associated w/ response.
 - By removing term (setting $\hat{\beta}_i = 0$), we can obtain a more interpretable model.
 - Note: Least squares will hardly ever result in $\hat{\beta}_i = 0$.

goal: need variable selection.

Same ideas apply to logistic regression.

1 Subset Selection

We consider methods for selecting subsets of predictors.

1.1 Best Subset Selection

To perform *best subset selection*, we fit a separate least squares regression for each possible combination of the p predictors.

$\binom{p}{k}$ models for each # of predictors in model (k).

Algorithm:

1. Let μ_0 denote null model - no predictors.
2. For $k=1, \dots, p$
 - (a) Fit all $\binom{p}{k}$ models that contain k predictors.
 - (b) Pick the best of those call it μ_k . "Best" defined by \downarrow RSS, \uparrow R^2
3. Select a single best model from $\mu_0, \mu_1, \dots, \mu_p$ using CV error, C_p , AIC/BIC, or adjusted R^2 more later.

We can't use R^2 for step 3. as $k \uparrow, R^2 \uparrow$ always.

Why might we not want to do this procedure at all?

We can perform something similar with logistic regression.

Fitting 2^p models! $p=10 \Rightarrow 1000$ models.

1.2 Stepwise Selection

For computational reasons, best subset selection cannot be performed for very large p .

impossible for $p \geq 40$

Best subset also suffers when p is large because w/ large search space

We can find good models on training data that perform poorly on test data.

high variability; overfitting of coeffs can occur.

Stepwise selection is a computationally efficient procedure that considers a much smaller subset of models.

Forward Stepwise Selection: start w/ no predictors and add predictors one at a time until all predictors are in the model. Choose the "best" from these.

1. Let μ_0 denote the null model - no predictors.
2. For $k=0, \dots, p-1$
 - (a) consider all $p-k$ models that augment predictors in μ_k w/ 1 additional predictor.
 - (b). Choose the best among $p-k$ and call it μ_{k+1} ($\uparrow R^2$, \downarrow RSS)
3. Select a single best model from μ_0, \dots, μ_p using CV error, C_p , AIC/BIC, or adjusted R^2 .

Now we are fitting $1 + \sum_{k=0}^{p-1} (p-k) = 1 + \frac{p(p+1)}{2}$ models!

Backward Stepwise Selection: Begin w/ full model and take predictors away one at a time until we get to null model.

1. Let M_p denote the full model - all p predictors
2. $k = p, p-1, \dots, 1$:
 - (a) Consider all k models that contain all but one of the predictors in M_k ($k-1$ predictors).
 - (b) Choose best among them and call it M_{k-1} ($\uparrow R^2, \downarrow RSS$).
3. Select single best model using CV error, etc.

* Neither forward nor backwards stepwise selection are guaranteed to find the best model containing a subset of the p predictors.

When $p > n$: forward selection can be used (but only up to $n-1$ predictors, not p).

1.3 Choosing the Optimal Model

Best subset, forward, backward select all need to pick "best" model - according to test error.

RSS & R^2 are proxies for training error \Rightarrow not good estimates of test error.
 ① estimate this directly
 ② adjust training for model size.

② $C_p = \frac{1}{n} (RSS + 2d\hat{\sigma}^2)$
 ↑ estimate of variance of ϵ full model
 # predictors in subset model

add penalty to training error (RSS) to adjust for underestimation of test error

as $d \uparrow, C_p \uparrow$

② AIC & BIC For maximum likelihood fits (≠ linear fits w/ least squares).

Choose model w/ lowest BIC

$AIC = \frac{1}{n} \hat{\sigma}^2 (RSS + 2d\hat{\sigma}^2)$

$BIC = \frac{1}{n} \hat{\sigma}^2 (RSS + \log(n)d\hat{\sigma}^2)$

since for $n \gg 7 \Rightarrow \log(n) > 2 \Rightarrow BIC$ is heavier penalty for adding variables \Rightarrow results in smaller models.

② Adjusted R^2 (least squares models)

$R^2 = 1 - \frac{RSS}{TSS}$ always \uparrow as $d \uparrow$

$Adj R^2 = 1 - \frac{RSS/(n-d-1)}{TSS/(n-1)}$ [choose model w/ highest $Adj R^2$]

① Validation and Cross-Validation

- Directly estimate test error w/ validation or CV and choose model w/ lowest estimated error.

\Rightarrow very general (can be used w/ any model) even when it's not clear how many "predictors" are in the model.

Now we have fast computers \Rightarrow CV is preferred.

proportional \Rightarrow same answer

2 Shrinkage Methods

The subset selection methods involve using least squares to fit a linear model that contains a subset of the predictors. As an alternative, we can fit a model with all p predictors using a technique that constrains (*regularizes*) the estimates.

↳ shrink towards zero

Shrinking the coefficient estimates can significantly reduce their variance!

Helps us avoid overfitting.

2.1 Ridge Regression

Recall that the least squares fitting procedure estimates β_1, \dots, β_p using values that minimize

“residual sum of squares”

$$RSS = \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2$$

Ridge Regression is similar to least squares, except that the coefficients are estimated by minimizing

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda \sum_{j=1}^p \beta_j^2 = RSS + \lambda \sum_{j=1}^p \beta_j^2$$

$\hat{\beta}^R$

note we do not penalize β_0
we want to penalize relationships
not the intercept (mean response
when $x_{i1} = \dots = x_{ip} = 0$)

$\lambda \geq 0$ “how much to penalize magnitude of coeffs”
tuning parameter (determined separately from fitting procedure).

trades off 2 criteria: minimize RSS to fit data well

$\lambda \sum_{j=1}^p \beta_j^2$ shrinkage penalty small when β_j close to zero \Rightarrow shrinks estimates towards zero.

The tuning parameter λ serves to control the impact on the regression parameters.

when $\lambda=0$ penalty has no effect and ridge regression = least squares.

As $\lambda \rightarrow \infty$, impact of penalty grows $\hat{\beta}^R \rightarrow 0$

Ridge regression will produce a different set of coefficients for each penalty $(\hat{\beta}_\lambda^R)$

selecting a good λ is critical! How to choose? CV.

The standard least squares coefficient estimates are scale invariant.

Multiply X_j by a constant c leads to a scaling of least squares coef by a factor of $1/c$.
 \Rightarrow regardless of how j^{th} predictor is scaled, $X_j \hat{\beta}_j$ will remain the same.

In contrast, the ridge regression coefficients $\hat{\beta}_\lambda^R$ can change substantially when multiplying a given predictor by a constant.

e.g. say we have an income variable in ① dollars and ② thousands of dollars.
 ① = 1000 × ②

due to the sum of squared coef term, this change does not simply result in the coefficient estimate to change by a factor of 1000.

$\Rightarrow X_j \hat{\beta}_{j\lambda}^R$ depends not only on λ but also on the scaling of X_j
 (may even depend on scaling of other predictors!).

Therefore, it is best to apply ridge regression after standardizing the predictors so that they are on the same scale:

i.e. standard deviation of one.

$$\tilde{x}_{ij} = \frac{x_{ij}}{\sqrt{\frac{1}{n} \sum_{i=1}^n (x_{ij} - \bar{x}_j)^2}}$$

estimate of
st. dev. of j^{th} predictor.

- ① standardize data
- ①.5 tune model to choose λ (w/ CV)
- ② fit ridge regression w/ chosen λ

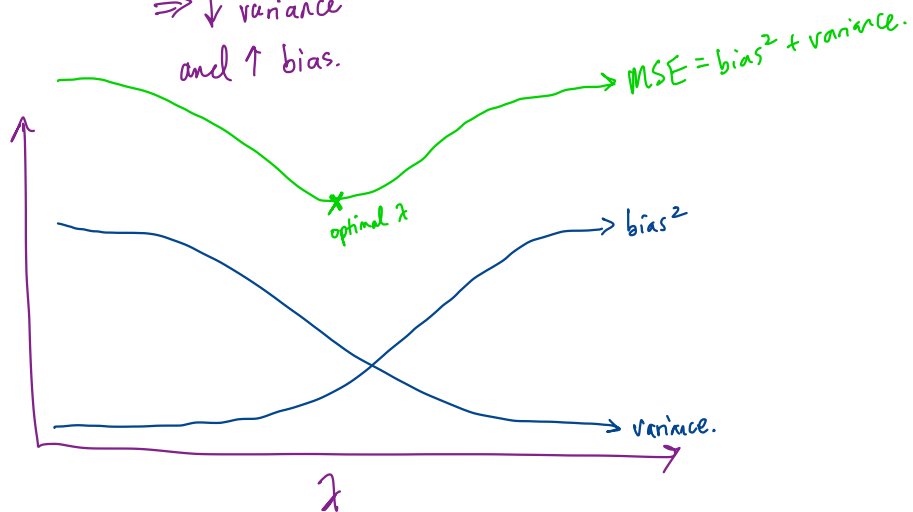
best
practice

Why does ridge regression work?

Because of the bias-variance trade-off!

As $\lambda \uparrow$: the flexibility of the ridge regression fit \downarrow

$\Rightarrow \downarrow$ variance
and \uparrow bias.



In situations where relationship between response and predictors \approx linear,
least squares will have low bias in its estimates

When p almost as large as $n \rightarrow$ least squares has variability!

(if $p > n$ least squares doesn't have a solution!)

\rightarrow ridge regression can still perform well in these scenarios by trading off a small amount of bias for a decrease in variance.

\Rightarrow ridge regression works best in high variance scenarios.

Also

Cost advantage of ridge regression over subset selection

b/c for a fixed λ , only fitting one model! (very fast model to fit).

Ridge regression improves predictive performance.

Does it also help us with interpretation? No!

2.2 The Lasso

Ridge regression does have one obvious disadvantage.

Unlike best subset, forward and backward selection ridge regression will include all p variables in final model.

penalty $\lambda \sum \beta_j^2$ will shrink all $\beta_j \rightarrow 0$ but $\beta_j \neq 0$ (unless $\lambda = \infty$).

This may not be a problem for prediction accuracy, but it could be a challenge for model interpretation when p is very large.

We will always have all variables in model, whether there is a true relationship w/ Y or not!

The lasso is an alternative that overcomes this disadvantage. The lasso coefficients $\hat{\beta}_\lambda^L$ minimize

As with ridge regression, the lasso shrinks the coefficient estimates towards zero.

As a result, lasso models are generally easier to interpret.

Why does the lasso result in estimates that are exactly equal to zero but ridge regression does not? One can show that the lasso and ridge regression coefficient estimates solve the following problems

In other words, when we perform the lasso we are trying to find the set of coefficient estimates that lead to the smallest RSS, subject to the constraint that there is a budget s for how large $\sum_{j=1}^p |\beta_j|$ can be.

2.3 Tuning

We still need a mechanism by which we can determine which of the models under consideration is “best”.

For both the lasso and ridge regression, we need to select λ (or the budget s).

How?