## **Chapter 7: Moving Beyond Linarity**

So far we have mainly focused on linear models.

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Jinear models are relatively simple to describe and implement. +: interpret & inference -: can have limited predictive performance because linearity assumption is alway an approximation (may not be a good one).

Previously, we have seen we can improve upon least squares using ride regression, the lasso, principal components regression, and more.

improvenat obtailed by reducing complexity of linear modules => lover variance of estimates. still a linear model! Can only be improved so much.

Through simple and more sophisticated extensions of the linear model, we can relax the linearity assumption while still maintiaining as much interpretability as possible.

<u>Polynomial regression</u>: adding extra predictors that are original variables raised to a power e.g. oubic regression use X, X<sup>2</sup>, X<sup>3</sup> as predictors, e.g. Y = β<sub>0</sub> + β<sub>1</sub>X<sub>T</sub> β<sub>2</sub>X<sup>2</sup> + β<sub>3</sub>X<sup>3</sup> + E
 +: non-linear fit
 -: with large powers, polynomial can take very strange shapes (cspecially at doundary).
 (2) Step functions : cut the range of predictor into K district regions (to produce categorical variable). Fit a piecewise constant function to (binned) X.

3 Regression Splines : more flexible than polynomials & step functions (extends both) idea: cut range of X into K district regions & polynomial is fit within each region polynomials constrained so they smoothy joiled.

(4) Generalized additive models: extend above ideas to deal u/ multiple predictors.

Note: We can talk about regression or classification, e.g. Logistic regression (polynomial): P(Y=11X) = (porp(porp(x+...+paxd)) (texp(porp(x+...+paxd)) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(porp(x+...+paxd))) (texp(p

## **1** Step Functions

Using polynomial functions of the features as predictors imposes a *global* structure on the non-linear function of X.

We can instead use *step-functions* to avoid imposing a global structure.

idea: Break range of X into bias and fit ditterat constant b each bia. letail: (1) create cut points  $c_{13-3}c_{K}$  in the may of X (2) construct K+1 new variables.  $c_{0}(X) = II(X < c_{1})$   $c_{1}(X) = II(c_{1} \le X < c_{2})$   $\vdots$   $c_{K-1}(X) = II(c_{K-1} \le X < c_{K})$   $c_{K}(X) = II(c_{K} \le X)$ (3) Use least squares to fit a linear modul using  $c_{1}(X)$ ,  $c_{2}(X)$ , -3  $c_{K}(X)$   $Y = B_{0} + \beta_{1}C_{1}(X) + ... + \beta_{K}C_{K}(X) + \Sigma$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = \{0 \ o.Y.\}$  indicator variable $<math>c_{0}(X) = II(X < c_{1}) = II$ 

For a given value of X, at most one of  $C_1, \ldots, C_K$  can be non-zero.

Co(X) + C(X) + ... + CK(X) = I since X must be in exactly one interval.

When 
$$X < c_1 \Rightarrow$$
 all of predictors  $C_{1,-1}, C_K = O$   
 $\Rightarrow$  Bo interpreted as the mean value of  $Y$  when  $X < C_1$   
B; represent the average increase in the response for  $C_j \leq X < C_{j+1}$  relative  
to  $X < C_1$ .

We can also fit a logistic repression model for classification.

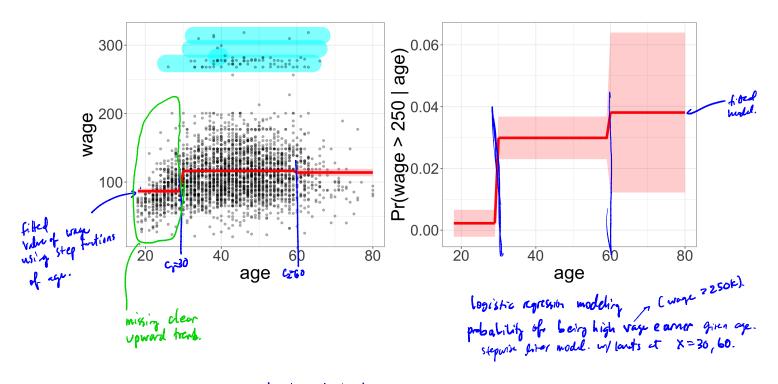
$$P(Y=|X) = \frac{e \times p(\beta_0 + \beta_1 C_1(X) + \dots + \beta_k C_k(X))}{1 + e \times p(\beta_0 + \beta_1 C_1(X) + \dots + \beta_k C_k(X))}$$

 $\mathbf{2}$ 

| yeara  | ge | maritl                      | race        | edu-<br>cation       |                                 | job-<br>class               | health               | health_ins | logwage  | wage      |
|--------|----|-----------------------------|-------------|----------------------|---------------------------------|-----------------------------|----------------------|------------|----------|-----------|
| 2006 1 | 10 | 1.<br>Never<br>Mar-<br>ried | 1.<br>White | 1. <<br>HS<br>Grad   | 2. Mid-<br>dle<br>At-<br>lantic | 1.<br>Indus-<br>trial       | 1.<br><=Good         | 2. No      | 4.318063 | 75.04315  |
| 2004 2 | 24 | 1.<br>Never<br>Mar-<br>ried | 1.<br>White | Col-<br>lege<br>Grad | At-<br>lantic                   | Infor-<br>ma-<br>tion       | 2.<br>>=Very<br>Good | 2. No      | 4.255273 | 70.47602  |
| 2003 4 |    | 2.<br>Mar-<br>ried          | 1.<br>White | Some                 | 2. Mid-<br>dle<br>At-<br>lantic | 1.<br>Indus-<br>trial       | 1.<br><=Good         | 1. Yes     | 4.875061 | 130.98218 |
| 2003 4 | 13 | 2.<br>Mar-<br>ried          | 3.<br>Asian | Col-<br>lege         | 2. Mid-<br>dle<br>At-<br>lantic | 2.<br>Infor-<br>ma-<br>tion | 2.<br>>=Very<br>Good | 1. Yes     | 5.041393 | 154.68529 |

Example: Wage data. for agroup of 3000 mile workers in mid-otherfic begion

 $c_1 = 30$  $c_2 = 60$ 



Unless there are natural breakpoints in the predictor, piecewise constant can miss trends.

## **2** Basis Functions

Polynomial and piecewise-constant regression models are in face special cases of a *basis function approach*.

Idea:

Instead of fitting the linear model in X, we fit the model

$$\gamma_i = \beta_0 + \beta_1 b_1(\chi_i) + \dots + \beta_K b_K(\chi_i) + \varepsilon_i$$

Note that the basis functions are fixed and known. We choose them ahead of the.

ex: polynomial regression 
$$b_j(x_i) = x_i^j$$
,  $j = 1, ..., d$ .

We can think of this model as a standard linear model with predictors defined by the basis functions and use least squares to estimate the unknown regression coefficients.

>

## **3** Regression Splines

*Regression splines* are a very common choice for basis function because they are quite flexible, but still interpretable. Regression splines extend upon polynomial regression and piecewise constant approaches seen previously.

# start

### **3.1** Piecewise Polynomials

Instead of fitting a high degree polynomial over the entire range of X, piecewise polynomial regression involves fitting separate low-degree polynomials over different regions of X.

e.g. one kot at C ammunu x fit two polynomials to the data one on s-but for XZC one an s-but for XZC each polynomial can be for using least squares. For example, a piercewise cubic with no knots is just a standard cubic polynomial.

A pieacewise cubic with a single knot at point c takes the form

$$\gamma_{i} = \begin{cases} P_{11} + \beta_{12} \chi_{i}^{2} + \beta_{21} \chi_{i}^{2} + \beta_{31} \chi_{i}^{3} + \varepsilon_{i} & iF \quad \chi_{i} < c \\ \beta_{01} + \beta_{12} \chi_{i} + \beta_{22} \chi_{i}^{2} + \beta_{32} \chi_{i}^{3} + \varepsilon_{i} & iF \quad \chi \geq c \end{cases}$$

Using more knots leads to a more flexible piecewise polynomial.

If we place k busts => fit k+1 polynomials

In general, we place K knots throughout the range of X and fit K + 1 polynomial regression models.

#### **3.2** Constraints and Splines

To avoid having too much flexibility, we can *constrain* the piecewise polynomial so that the fitted curve must be continuous.

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i.e. there cannot be a jump at knots.
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To go further, we could add two more constraints

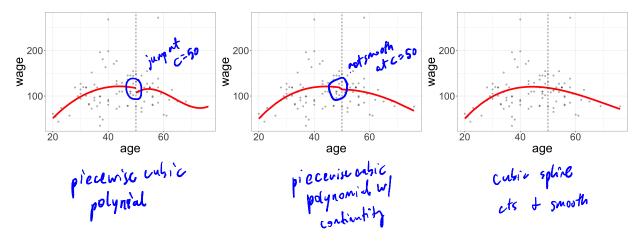
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1) 1st derivative of piecewise polynomial must be contributous
2) Ind derivative of precessive polynomial must be cts.
```

In other words, we are requiring the piecewise polynomials to be *smooth*.

Each constraint that we impose on the piecewise cubic polynomials effectively frees up one degree of freedom, by reducing the complexity of the resulting fit.

The fit with continuity and 2 smoothness contraints is called a *spline*.

A degree-d spline is a piece wise degree-d polynomial u/ continuity in servicious up to degree d-1 at each prot.



#### 3.3 Spline Basis Representation

Fitting the spline regression model is more complex than the piecewise polynomial regression. We need to fit a degree d piecewise polynomial and also constrain it and it's d - 1 derivatives to be continuous at the knots.

We can use the basis model to represent a regression spline.

eq:  

$$y_i = \beta_0 + \beta_1 \beta_1(x_i) + \beta_2 \beta_2(x_i) + \dots \beta_{K+3} \beta_{K+3}(x_i) + \varepsilon_i$$
  
so the most direct way to represent a cubic spline is to start with the basis for a cubic poly-  $\rightarrow \int_{a}^{b} \int_{a}^{b} \int_{a}^{b} \int_{a}^{c} \int_{a}^{$ 

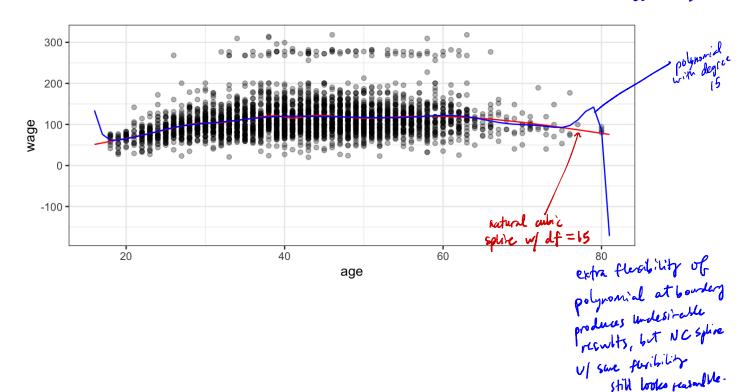
additional constraint produces more stable astimutes at the boundaries.

#### **3.4** Choosing the Knots

When we fit a spline, where should we place the knots? > tegression spline is most flexible in regions that contain a lot of boots (bootsing rapidly). > place boots where we think the function will vary rapidly and less mots when findin is studie. More common in practice : place them unfirmly. To place knots: choose desired degrees of findom (flexibility) & use software to automatically place # knots at uniform grantiles of deta. How many knots should we use? Lish how many df should be use? Use CV! Use k giving smallert CV MSE (CV error).

#### **3.5** Comparison to Polynomial Regression

Regression splines often give superior results to polynomial regression. >> Polynomial regression must use high degrees to achieve flexibility (e.g. X<sup>15</sup>), but regression splins introduce flexibility through boots (fixed degree polynomials) => more stability (csp. of Lourden'ss)



## **4** Generalized Additive Models

So far we have talked about flexible ways to predict Y based on a single predictor X.

These approaches can be seen as extensions of simple linear regression wodl.  $Y = \beta_0 + \beta_1 \otimes + \epsilon$ extension: basis footons  $f \times$ 

*Generalized Additive Models (GAMs)* provide a general framework for extending a standard linear regression model by allowing non-linear functions of each of the variables while maintaining *additivity*.

flexibly predict Y on General predictors X1, ,-, Xp. still additive models 4.1 GAMs for Regression can be used for regression or classification.

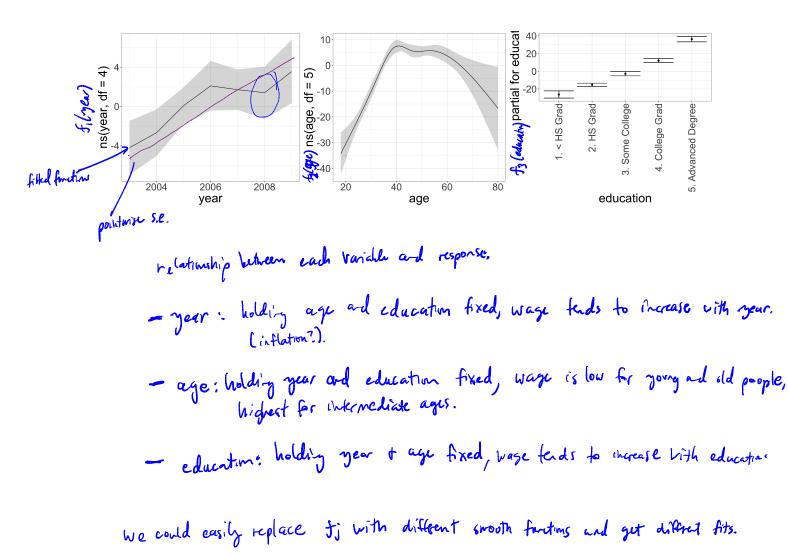
A natural way to extend the multiple linear regression model to allow for non-linear relationships between feature and response:

linear regression : 
$$\mathcal{J}_{i} = \beta_{0} + \beta_{i} x_{ii} + \beta_{2} x_{2i} + \dots + \beta_{p} x_{pi} + \varepsilon_{i}$$
  
extension idea! replace each linear temponent  $\beta_{j} x_{ij}$  with a smooth non-linear function  
 $\Longrightarrow GAM$ :  $\mathcal{Y}_{i} = \beta_{0} + \sum_{j=1}^{p} f_{j}(x_{ij}) + \varepsilon_{i}$   
 $= \beta_{0} + f_{1}(x_{i}) + f_{2}(x_{i2}) + \dots + f_{p}(x_{ip}) + \varepsilon_{i}$   
"additive" because we calculate asequate for the each predictor  $X_{j}$  and add then  
together.  
possibilities for  $f_{j}$ :  
-linear component (leads + linear regression).

The beauty of GAMs is that we can use our fitting ideas in this chapter as building blocks for fitting an additive model.

Example: Consider the Wage data. quartitutive categorical.

Wage = 
$$\beta_0 + f_1(year) + f_2(age) + f_3(education) + \epsilon$$
  
where  $f_1$  is natural splite  $w/ 4 df$   
 $f_2$  is natural cubic splite  $w/ 5 df$   
 $f_3$  constant functions for each value (dummy variables)



just read to change basis & use least squares.

#### Pros and Cons of GAMs

#### Advantages

For fully general moduls, he have to look for even more flexible approaches like random forests or boosting (next week!) GRAMS provide a useful compromise between Linear and respiranetic approacher.

#### 4.2 GAMs for Classification



GAMs can also be used in situations where Y is categorical. Recall the logistic regression model:

A natural way to extend this model is for non-linear relationships to be used.

$$\log\left(\frac{\rho(x)}{1-\rho(x)}\right) = \beta_0 + f_1(x_1) + \dots + f_p(x_p)$$
  
logistic regression GAM

Example: Consider the Wage data.

